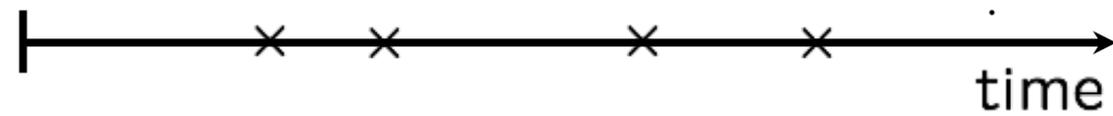


## LECTURE 22: The Poisson process

- Definition of the Poisson process
  - applications
- Distribution of number of arrivals
- The time of the  $k$ th arrival
- Memorylessness
- Distribution of interarrival times

## Definition of the Poisson process

### Poisson



- Numbers of arrivals in disjoint time intervals are **independent**

$P(k, \tau)$  = Prob. of  $k$  arrivals in interval of duration  $\tau$

- **Small interval probabilities:**

For VERY small  $\delta$ :

$$P(k, \delta) \approx \begin{cases} 1 - \lambda\delta & \text{if } k = 0 \\ \lambda\delta & \text{if } k = 1 \\ 0 & \text{if } k > 1 \end{cases} \quad P(k, \delta) = \begin{cases} 1 - \lambda\delta + O(\delta^2) & \text{if } k = 0 \\ \lambda\delta + O(\delta^2) & \text{if } k = 1 \\ 0 + O(\delta^2) & \text{if } k > 1 \end{cases}$$

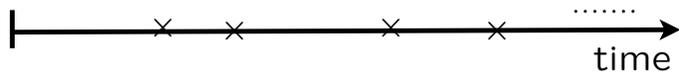
### Bernoulli



- Independence
- **Time homogeneity:**  
Constant  $p$  at each slot

$\lambda$ : "arrival rate"

## Applications of the Poisson process



- Deaths from horse kicks in the Prussian army (1898)
- Particle emissions and radioactive decay
- Photon arrivals from a weak source
- Financial market shocks
- Placement of phone calls, service requests, etc.



Siméon Denis Poisson  
(1781-1840)

(This image is in the public domain.  
Source: [Wikipedia](#))

## The Poisson PMF for the number of arrivals



- $N_\tau$ : arrivals in  $[0, \tau]$        $P(k, \tau) = \mathbf{P}(N_\tau = k)$

$n = \tau/\delta$  intervals/slots of length  $\delta$

$\mathbf{P}$ (some slot contains two or more arrivals)

$$N_\tau \approx \text{binomial} \quad p = \lambda\delta + O(\delta^2)$$

$$np =$$

### Bernoulli

$$p_S(k) = \frac{n!}{(n-k)!k!} \cdot p^k(1-p)^{n-k},$$

$$k = 0, \dots, n$$

$$\lambda = np \quad n \rightarrow \infty \quad p \rightarrow 0$$

For fixed  $k = 0, 1, \dots,$

$$p_S(k) \rightarrow \frac{\lambda^k}{k!} e^{-\lambda},$$

$$P(k, \tau) = \frac{(\lambda\tau)^k e^{-\lambda\tau}}{k!}, \quad k = 0, 1, \dots$$

## Mean and variance of the number of arrivals

$$P(k, \tau) = \mathbf{P}(N_\tau = k) = \frac{(\lambda\tau)^k e^{-\lambda\tau}}{k!}, \quad k = 0, 1, \dots$$

$$\mathbf{E}[N_\tau] = \sum_{k=0}^{\infty} k \frac{(\lambda\tau)^k e^{-\lambda\tau}}{k!} = \dots$$

$$N_\tau \approx \text{Binomial}(n, p)$$

$$n = \tau/\delta, \quad p = \lambda\delta + O(\delta^2)$$

$$\mathbf{E}[N_\tau] = \lambda\tau$$

$$\text{var}(N_\tau) = \lambda\tau$$

## Example

- You get email according to a Poisson process, at a rate of  $\lambda = 5$  messages per hour.
- Mean and variance of mails received during a day =
- $P(\text{one new message in the next hour}) =$

$$P(k, \tau) = \frac{(\lambda\tau)^k e^{-\lambda\tau}}{k!}, \quad k = 0, 1, \dots$$

- $P(\text{exactly two messages during each of the next three hours}) =$

$$E[N_\tau] = \lambda\tau$$

$$\text{var}(N_\tau) = \lambda\tau$$

## The time $T_1$ until the first arrival

$$P(k, \tau) = \frac{(\lambda\tau)^k e^{-\lambda\tau}}{k!}, \quad k = 0, 1, \dots$$

- Find the CDF:  $P(T_1 \leq t) =$

$$f_{T_1}(t) = \lambda e^{-\lambda t}, \quad \text{for } t \geq 0$$

Exponential( $\lambda$ )

**Memorylessness:** conditioned on  $T_1 > t$ ,  
the PDF of  $T_1 - t$  is again exponential

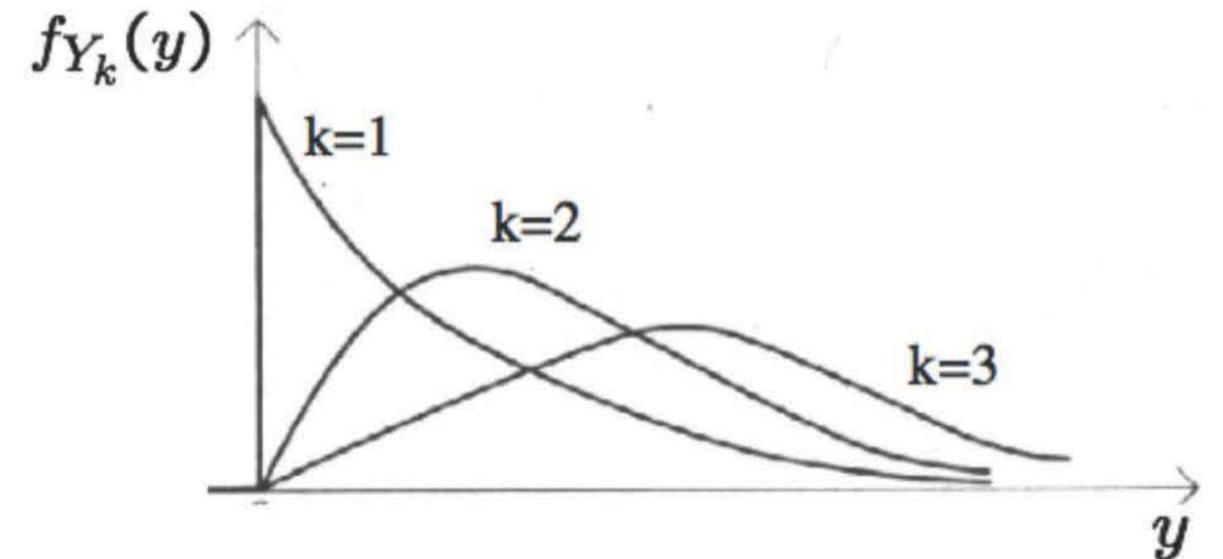
## The time $Y_k$ of the $k$ th arrival

$$P(k, \tau) = \frac{(\lambda\tau)^k e^{-\lambda\tau}}{k!}, \quad k = 0, 1, \dots$$

- Can derive its PDF by first finding the CDF
- More intuitive argument:

$$f_{Y_k}(y) \delta \approx \mathbf{P}(y \leq Y_k \leq y + \delta) =$$

$$\text{Erlang distribution: } f_{Y_k}(y) = \frac{\lambda^k y^{k-1} e^{-\lambda y}}{(k-1)!}, \quad y \geq 0$$

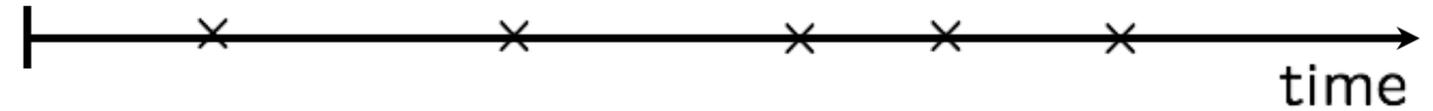


## Memorylessness and the fresh-start property

- Analogous to the properties for the Bernoulli process
  - plausible, given the relation between the two processes
  - use intuitive reasoning
  - can be proved rigorously

## Memorylessness and the fresh-start property

- If we start watching at time  $t$ ,



we see Poisson process, independent of the history until time  $t$

– time until next arrival:

- If we start watching at time  $T_1$ ,

we see Poisson process, independent of the history until time  $T_1$

– hence: time between first and second arrival,  $T_2 = Y_2 - Y_1$  is:

– similarly for all  $T_k = Y_k - Y_{k-1}$ ,  $k \geq 2$

$Y_k = T_1 + \dots + T_k$  is sum of i.i.d. exponentials

$$\mathbf{E}[Y_k] = k/\lambda \quad \mathbf{var}(Y_k) = k/\lambda^2$$

- An equivalent definition
- A simulation method

## Bernoulli/Poisson relation



$$n = \tau/\delta,$$

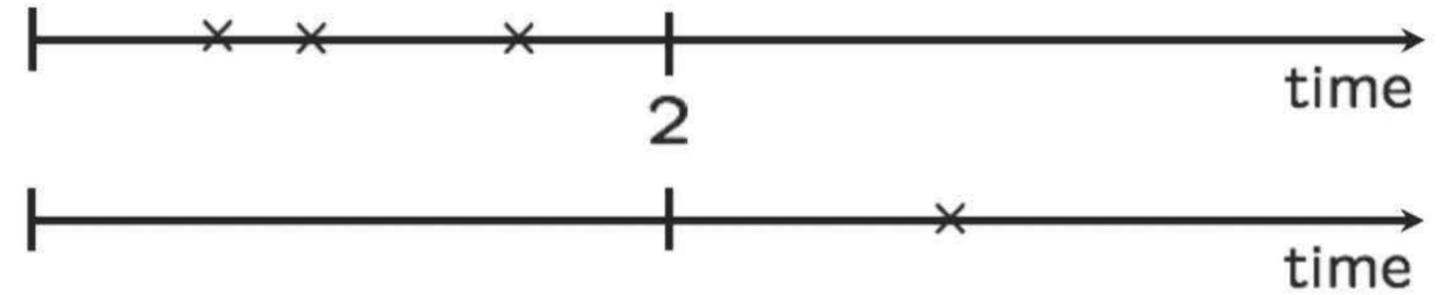
$$p = \lambda\delta$$

$$np = \lambda\tau$$

	<b>POISSON</b>	<b>BERNOULLI</b>
Times of Arrival	Continuous	Discrete
Arrival Rate	$\lambda$ /unit time	$p$ /per trial
PMF of # of Arrivals	Poisson	Binomial
Interarrival Time Distr.	Exponential	Geometric
Time to $k$ -th arrival	Erlang	Pascal

## Example: Poisson fishing

- Fish are caught as a Poisson process,  $\lambda = 0.6/\text{hour}$ 
  - fish for two hours;
  - if you caught at least one fish, stop
  - else continue until first fish is caught



$P(\text{fish for more than two hours}) =$

$P(\text{fish for more than two and less than five hours}) =$

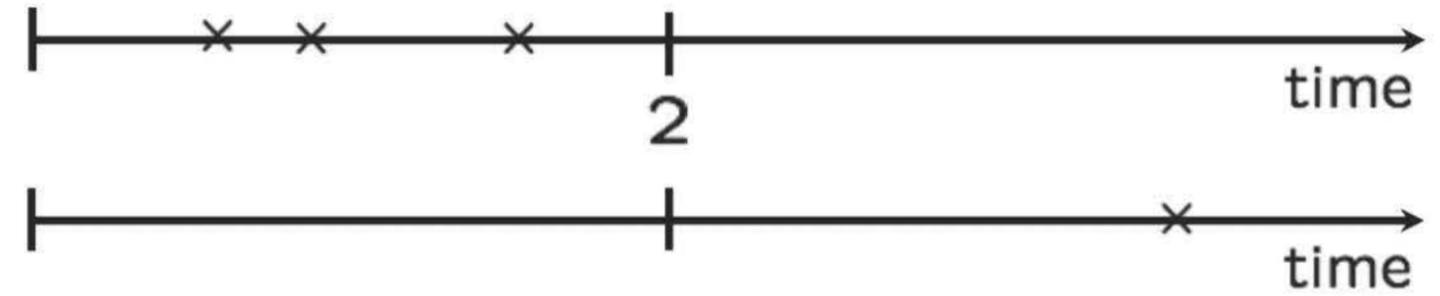
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$$E[N_\tau] = \lambda\tau$$

$$f_{Y_k}(y) = \frac{\lambda^k y^{k-1} e^{-\lambda y}}{(k-1)!}$$

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$P(\text{catch at least two fish}) =$

$E[\text{future fishing time} \mid \text{already fished for three hours}] =$

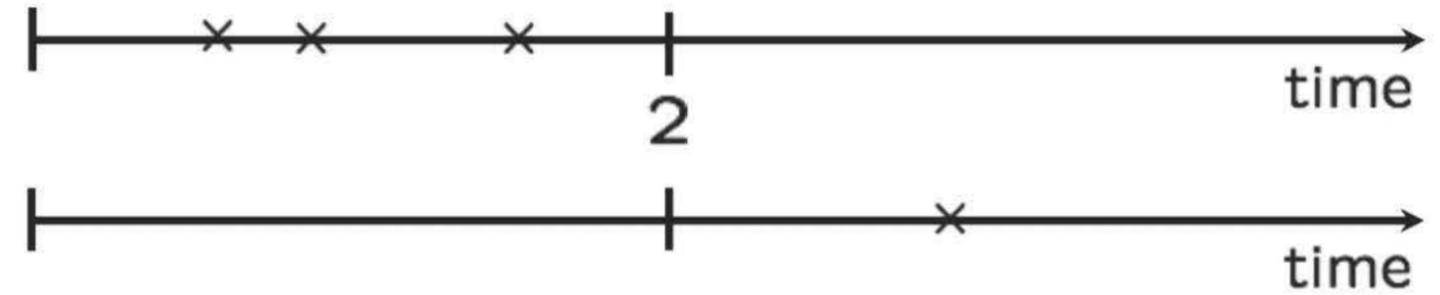
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$\mathbf{E}[\text{total fishing time}] =$

$\mathbf{E}[\text{number of fish}] =$

$$P(k, \tau) = \frac{(\lambda\tau)^k e^{-\lambda\tau}}{k!}$$

$$\mathbf{E}[N_\tau] = \lambda\tau$$

$$f_{Y_k}(y) = \frac{\lambda^k y^{k-1} e^{-\lambda y}}{(k-1)!}$$

MIT OpenCourseWare  
<https://ocw.mit.edu>

Resource: Introduction to Probability  
John Tsitsiklis and Patrick Jaillet

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